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Probability Distribution of Curvatures of Isosurfaces in Gaussian Random Fields

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Abstract

An expression for the joint probability distribution of the principal curvatures at an arbitrary point in the ensemble of isosurfaces defined on isotropic Gaussian random fields on \mathbb{R}^n is derived. The result is obtained by deriving symmetry properties of the ensemble of second derivative matrices of isotropic Gaussian random fields akin to those of the Gaussian orthogonal ensemble.

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I. INTRODUCTION

We closely follow the notation in [1] and [2]. Let \mathbf{T} be a tensor with rank a and dimensions (d_1, \dots, d_a) . The bijective linear map vec associates \mathbf{T} to the vector $\text{vec } \mathbf{T}$ in \mathbb{R}^N , $N = \prod_{i=1}^a d_i$, with entry $(\text{vec } \mathbf{T})_k$, $k \in \{1, \dots, \prod_{i=1}^a d_i\}$, given by $(\text{vec } \mathbf{T})_k = \mathbf{T}_{i_1, \dots, i_a}$ with $i_j \in \{1, \dots, d_j\}$ uniquely defined by $k = 1 + \sum_{j=1}^a (i_j - 1) \prod_{k=1}^{j-1} d_k$. Let $d_i = n$, $i = 1, \dots, a$. The linear operator diag maps \mathbf{T} to the vector $\text{diag } \mathbf{T}$ in \mathbb{R}^n with entry $(\text{diag } \mathbf{T})_k$, $k \in \{1, \dots, n\}$, given by $\mathbf{T}_{k,k,\dots,k}$. If restricted to the set D of diagonal matrices, $D \ni \mathbf{D} \rightarrow \text{diag } \mathbf{D}$ is bijective, and therefore the inverse mapping diag^{-1} is well-defined. Let $a = 2$ and $d_1 = d_2 = n$. The linear operator uni maps \mathbf{T} to the vector $\text{uni } \mathbf{T} \in \mathbb{R}^{n(n+1)/2}$ with entry $(\text{vec } \mathbf{T})_k$, $k \in \{1, \dots, n(n+1)/2\}$, given by $(\text{vec } \mathbf{T})_k = \mathbf{T}_{i,j}$ uniquely defined by $k = (j-1)n - j(j-1)/2 + i$. This operator maps a matrix \mathbf{T} to a vector containing the entries of \mathbf{T} read along its columns, but ignoring the elements above the main diagonal.

Following the notation in [1], the (n, n) identity matrix is denoted by \mathbf{I}_n , \mathbf{C}_n denotes the (n^2, n^2) commutation matrix, i.e., the unique (n^2, n^2) matrix such that $\text{vec } \mathbf{T}^\top = \mathbf{C}_n \text{vec } \mathbf{T}$ for all (n, n) matrices \mathbf{T} . The matrix \mathbf{D}_n denotes the $(n^2, n(n+1)/2)$ duplication matrix, i.e., the unique $(n^2, n(n+1)/2)$ matrix such that $\text{vec } \mathbf{T} = \mathbf{D}_n \text{uni } \mathbf{T}$ for all (n, n) symmetric matrices \mathbf{T} . From this definition, we have $\text{uni } \mathbf{T} = \mathbf{D}_n^+ \text{vec } \mathbf{T}$, where \mathbf{A}^+ indicates the Moore-Penrose inverse of the matrix \mathbf{A} . The duplication matrix and the commutation matrix are related through the identity $\mathbf{D}_n \mathbf{D}_n^+ = (1/2)(\mathbf{I}_{n^2} + \mathbf{C}_n)$. Finally, $\mathbf{1}_{n,m}$ denotes an (n, m) matrix of ones.

A *random scalar field* in the set X is a stochastic process, i.e., an indexed set $\mathcal{F}_X = \{\mathcal{F}_x, x \in X\}$ of random variables \mathcal{F}_x defined over the same probability space $(\Omega, \sigma_\Omega, P)$. A *random tensor fields* is defined analogously, with \mathcal{F}_x a vector-valued function.

Let \mathcal{F}_X^1 and \mathcal{F}_X^2 be two random scalar fields as above, and assume that \mathcal{F}_x^i is zero mean, which, for the purposes of this work, implies no loss of generality. If the expectation $E\{\mathcal{F}_x^1 \mathcal{F}_y^2\}$ taken over the joint distribution of \mathcal{F}_X^1 and \mathcal{F}_X^2 is defined for all x and y in X , the function $R_{\mathcal{F}^1, \mathcal{F}^2}(x, y) = E\{\mathcal{F}_x^1 \mathcal{F}_y^2\}$ defines the *cross-covariance function* of the random fields. If $\mathcal{F}_X^1 = \mathcal{F}_X^2 = \mathcal{F}_X$, the notation is simplified to $R_{\mathcal{F}} \triangleq R_{\mathcal{F}, \mathcal{F}}$, and the function $R_{\mathcal{F}}$ is referred to as the *autocovariance function* of \mathcal{F}_X . The *conditional autocovariance function* of \mathcal{F}_X^1 given \mathcal{F}_X^2 , $R_{\mathcal{F}^1 | \mathcal{F}^2}$, is defined by taking the expectation of \mathcal{F}_x^1 over the conditional distribution of \mathcal{F}_x^1 and \mathcal{F}_y^1 given \mathcal{F}_X^2 . In the case of a random tensor field the definitions are analogous, with the product $\mathcal{F}_x^1 \mathcal{F}_y^2$ replaced by a tensor product $\text{vec } \mathcal{F}_x^1 \otimes (\text{vec } \mathcal{F}_y^2)^\top$ and the expectation taken over each entry of the tensor. Henceforth, the term “random field” will be used in reference to both scalar and tensorial random fields.

Let X be a vector space. Zero-mean random fields \mathcal{F}_X^1 and \mathcal{F}_X^2 on X are *wide-sense stationary* if their cross-covariance function $R_{\mathcal{F}_X^1, \mathcal{F}_X^2}(\mathbf{x}, \mathbf{y})$ satisfies $R_{\mathcal{F}_X^1, \mathcal{F}_X^2}(\mathbf{x}, \mathbf{y}) = R_{\mathcal{F}_X}(\mathbf{s})$ with $\mathbf{s} = \mathbf{x} - \mathbf{y}$ for all \mathbf{x} and \mathbf{y} in X . Henceforth we will assume that $X = \mathbb{R}^n$, and therefore the subscript X in \mathcal{F}_X can then be safely omitted by defining $\mathcal{F} \triangleq \mathcal{F}_{\mathbb{R}^n}$. A stationary random field on \mathbb{R}^n is *isotropic* if its autocovariance function $R_{\mathcal{F}}(\mathbf{s})$ satisfies $R_{\mathcal{F}}(\mathbf{s}) = \sigma^2 \rho(\|\mathbf{s}\|)$, where ρ is a correlation function [2] and $\|\cdot\|$ is the standard Euclidean norm in \mathbb{R}^n .

II. DERIVATIVES OF ISOTROPIC RANDOM FIELDS

Great simplification is achieved in the derivations that follow if $\rho(\|\mathbf{s}\|)$ can be rewritten as $\rho(\|\mathbf{s}\|) = r(\|\mathbf{s}\|^2)/\sigma^2 \Leftrightarrow \rho(\sqrt{\|\mathbf{s}\|}) = r(\|\mathbf{s}\|)/\sigma^2$. For $\|\mathbf{s}\| > 0$ the smoothness of $r(\|\mathbf{s}\|)$ is contingent upon that of $\rho(\|\mathbf{s}\|)$. However, for $\|\mathbf{s}\| = 0$ the non-differentiability of $\sqrt{\|\mathbf{s}\|}$ could be a problem. This is not the case, as shown in appendix. The symbols $\rho_0^{(i)}$ and $r_0^{(i)}$ denotes the i -th derivative of $\rho(\|\mathbf{s}\|)$ and $r(\|\mathbf{s}\|)$ with respect to $s = \|\mathbf{s}\|$ at $s = 0$.

Let $f : \mathbb{R}^n \rightarrow \mathbb{R}$ be a scalar function. The symbol $\frac{\partial^{a+b} f}{\partial \mathbf{x}^{aT} \partial \mathbf{x}^b}$ is used to describe the matrix of dimensions (n^a, n^b) of partial derivatives of f , i.e,

$$\left(\frac{\partial^{a+b} f}{\partial \mathbf{x}^{aT} \partial \mathbf{x}^b} \right)_{A,B} \triangleq \frac{\partial^{a+b} f}{\partial x_{a_1} \dots \partial x_{a_a} \partial x_{b_1} \dots \partial x_{b_b}},$$

where $a_i \in \{1, 2, \dots, n\}$ and $b_j \in \{1, 2, \dots, n\}$ are uniquely defined by $A = 1 + \sum_{i=1}^n (a_i - 1)n^{i-1}$ and $B = 1 + \sum_{j=1}^n (b_j - 1)n^{j-1}$. Second differentiability of the autocovariance function $R_{\mathcal{F}}(\mathbf{x}, \mathbf{y})$ of a random field at the pair (\mathbf{x}, \mathbf{y}) implies *mean square differentiability* of the random field itself, as demonstrated in theorem 2.4 of [2]. If a mean-square differentiable random field is stationary, its derivatives will also be stationary. For \mathcal{F}^1 and \mathcal{F}^2 jointly stationary with cross-covariance $R_{\mathcal{F}^1, \mathcal{F}^2}(\mathbf{s})$, we define $R_{\mathcal{F}^1, \mathcal{F}^2}^0 \triangleq R_{\mathcal{F}^1, \mathcal{F}^2}(\mathbf{0})$.

The theorem that follows is central to this work:

Theorem 1. *Let \mathcal{F} be an isotropic random field on \mathbb{R}^n with autocovariance function $R_{\mathcal{F}}(s) = \sigma^2 \rho(\|s\|)$ where $\rho(s)$ is four-differentiable. Let $\partial \mathcal{F}$ and $\partial^2 \mathcal{F}$ be the tensor fields defined as*

$$\partial \mathcal{F} \triangleq \frac{\partial \mathcal{F}}{\partial \mathbf{x}^T} \quad \text{and} \quad \partial^2 \mathcal{F} \triangleq \frac{\partial^2 \mathcal{F}}{\partial \mathbf{x}^T \partial \mathbf{x}}.$$

Then

$$R_{\partial\mathcal{F}}^0 = -\sigma^2 \rho_0^{(2)} \mathbf{I}_n, \quad (1a)$$

$$R_{\partial\mathcal{F},\partial^2\mathcal{F}}^0 = \mathbf{O}_{n,n^2}, \quad (1b)$$

$$R_{\partial^2\mathcal{F}}^0 = \sigma^2 \rho_0^{(4)} (\mathbf{I}_{n^2} + \mathbf{C}_n + \text{vec } \mathbf{I}_n \text{ vec}^T \mathbf{I}_n). \quad (1c)$$

Proof. Using lemma 4 in appendix we write $\rho(\|\mathbf{s}\|) = r(\|\mathbf{s}\|^2)/\sigma^2$. From the four-differentiability of r and theorem 2.4 in [2] we have

$$R_{\partial\mathcal{F}}(s) = -\frac{\partial^2 r(\|\mathbf{s}\|^2)}{\partial s^T \partial s}, \quad (2a)$$

$$R_{\partial\mathcal{F},\partial^2\mathcal{F}}(s) = -\frac{\partial^3 r(\|\mathbf{s}\|^2)}{\partial s^T \partial s^2}, \quad (2b)$$

$$R_{\partial^2\mathcal{F}}(s) = \frac{\partial^4 r(\|\mathbf{s}\|^2)}{\partial s^{2T} \partial s^2}. \quad (2c)$$

The chain rule can be used to expand (2a)–(2c), and substituting the identities $\frac{\partial \mathbf{x}}{\partial \mathbf{x}} = \frac{\partial \mathbf{x}^T}{\partial \mathbf{x}^T} = \mathbf{I}_n$, $\frac{\partial(\mathbf{x}^T \otimes \mathbf{I}_n)}{\partial \mathbf{x}^T} = \mathbf{I}_n \otimes \mathbf{I}_n$ and $\frac{\partial(\mathbf{I}_n \otimes \mathbf{x}^T)}{\partial \mathbf{x}^T} = \mathbf{C}_n$ in the result produces

$$R_{\partial\mathcal{F}}(s) = -4r^{(2)}(\|\mathbf{s}\|^2)s \otimes s^T - 2r^{(1)}(\|\mathbf{s}\|^2)\mathbf{I}_n,$$

$$R_{\partial\mathcal{F},\partial^2\mathcal{F}}(s) = -4\{2r^{(3)}(\|\mathbf{s}\|^2)s \otimes s^T \otimes s^T +$$

$$r^{(2)}(\|\mathbf{s}\|^2)(s^T \otimes \mathbf{I}_n +$$

$$\mathbf{I}_n \otimes s^T + s \otimes \text{vec}^T \mathbf{I}_n)\},$$

$$R_{\partial^2\mathcal{F}}(s) = 4\{8r^{(4)}(\|\mathbf{s}\|^2)s \otimes s \otimes s^T \otimes s^T +$$

$$8r^{(3)}(\|\mathbf{s}\|^2)(\text{vec } \mathbf{I}_n \otimes s^T \otimes s^T +$$

$$2s \otimes (\mathbf{I}_n \otimes s^T + s^T \otimes \mathbf{I}_n) + s \otimes s \otimes \text{vec } \mathbf{I}_n^T +$$

$$r^{(2)}(\|\mathbf{s}\|^2)(\mathbf{I}_n \otimes \mathbf{I}_n + \mathbf{C}_n + \text{vec } \mathbf{I}_n \otimes \text{vec}^T \mathbf{I}_n)\}.$$

Making $s = \mathbf{0}$ completes the proof. \square

III. CURVATURES OF GAUSSIAN RANDOM FIELDS

Let $f(\mathbf{x})$ be a second-differentiable scalar function on \mathbb{R}^n . For each $\mathbf{x} \in \mathbb{R}^n$ for which $\partial f / \partial \mathbf{x} \neq \mathbf{0}$ we define the set $F_{\mathbf{x}}$ as $F_{\mathbf{x}} = \{\mathbf{y} \in \mathbb{R}^n \text{ such that } f(\mathbf{y}) = f(\mathbf{x}) \text{ and } \partial f / \partial \mathbf{y} \neq \mathbf{0}\}$. If $F_{\mathbf{x}} \neq \emptyset$, $F_{\mathbf{x}}$ is a $(n-1)$ -hypersurface in \mathbb{R}^n [3]. Let $\partial f \triangleq \partial f / \partial \mathbf{x}^T$ and $\partial^2 f \triangleq \partial^2 f / \partial \mathbf{x}^T \partial \mathbf{x}$. The *principal*

curvatures of the hypersurface $F_{\mathbf{x}}$ at \mathbf{x} are given by the set of eigenvalues κ obtained by solving the eigenproblem

$$-\left(\mathbb{I}_n - \frac{\partial f \otimes (\partial f)^T}{\|\partial f\|^2}\right) \frac{\partial^2 f}{\|\partial f\|} \Big|_{\mathbf{x}} \mathbf{v} = \kappa \mathbf{v}, \quad (4)$$

with $\mathbf{v} \in \mathbb{R}^n$, $\|\mathbf{v}\| = 1$ and $\mathbf{v}^T \partial f = 0$ [4, pg. 138]. Let $\{\mathbf{n}_i, i = 1, \dots, n-1\}$ be an orthonormal basis for the null-space of ∂f , i.e., $\mathbf{n}_i^T \partial f = 0$ and $\mathbf{n}_i^T \mathbf{n}_j = \delta_{ij}$, and let \mathbf{N} be the matrix $\mathbf{N} = [\mathbf{n}_1 \dots \mathbf{n}_{n-1}]$. The eigenproblem in (4) can be rewritten as

$$-\frac{\mathbf{N}^T (\partial^2 f) \mathbf{N}}{\|\partial f\|} \Big|_{\mathbf{x}} \mathbf{u} = \kappa \mathbf{u}, \quad (5)$$

with $\mathbf{u} \in \mathbb{R}^{n-1}$, $\|\mathbf{u}\| = 1$. Equation (5) is still valid if the function f is the realization $\mathcal{F}(\omega)$, $\omega \in \Omega$, of a mean-square second-differentiable scalar random field \mathcal{F} on \mathbb{R}^n . Therefore the random tensor field \mathcal{K} of curvatures of isotropic Gaussian random fields is implicitly defined at \mathbf{x} such that $\partial \mathcal{F}_{\mathbf{x}}(\omega) \neq \mathbf{0}$ by the solutions of the equation

$$-\frac{\mathcal{N}^T (\partial^2 \mathcal{F}) \mathcal{N}}{\|\partial \mathcal{F}\|} \mathcal{U} = \mathcal{K} \mathcal{U}, \quad (6)$$

where \mathcal{N} is a random tensor field satisfying $\mathcal{N}_{\mathbf{x}}^T \partial \mathcal{F}_{\mathbf{x}} = 0$, and $\mathcal{N}_{\mathbf{x}}^T \mathcal{N}_{\mathbf{x}} = \mathbb{I}_{n-1}$, and \mathcal{U} is the tensor field such that $\mathcal{U}_{\mathbf{x}}(\omega)$ are the eigenvectors associated to the eigenvalues $\mathcal{K}_{\mathbf{x}}(\omega)$.

Henceforth we assume that \mathcal{F} is an isotropic, second-differentiable *Gaussian random field*, which is defined simply as an isotropic random field for which the joint distribution of any finite set of random variables $\{\mathcal{F}_A\}$, $A \in \mathbb{R}^n$, is Gaussian. This assumption implies that the zero-mean random tensors $\partial \mathcal{F}$ and $\partial^2 \mathcal{F}$ are also Gaussian, and therefore $\partial \mathcal{F}_{\mathbf{x}}$ and $\partial^2 \mathcal{F}_{\mathbf{x}}$ are fully characterized by their covariance matrices, given by (1a) and (1c) in theorem 1. However, because $\partial^2 \mathcal{F}$ is symmetric, its probability density must be handled with care, since $R_{\partial^2 \mathcal{F}}^0$ is not invertible.

The following lemma is a trivial corollary of the theorems in [1, sec. 7].

Lemma 1. *Let \mathbf{A}_n be a (n, n) invertible matrix. Then*

$$(\mathbf{A}_n \otimes \mathbf{A}_n) R_{\partial^2 \mathcal{F}}^0 (\mathbf{A}_n^{-1} \otimes \mathbf{A}_n^{-1}) = R_{\partial^2 \mathcal{F}}^0. \quad (7)$$

Let $\partial^2 \mathcal{F}$ be as in theorem 1, and let \mathcal{R} be a (n, m) , $n \geq m$, orthonormal tensor field independent of $\partial^2 \mathcal{F}$, i.e., a random tensor field such that for all $\mathbf{x} \in \mathbb{R}^n$ any realization $\mathcal{R}(\omega)$ of \mathcal{R} satisfies $\mathcal{R}_{\mathbf{x}}^T(\omega) \mathcal{R}_{\mathbf{x}}(\omega) = \mathbb{I}_m$ and $\mathcal{R}_{\mathbf{x}}$ is independent of $\partial^2 \mathcal{F}_{\mathbf{x}}$. Define $\partial^2 \mathcal{F}' \triangleq \mathcal{R}^T \partial^2 \mathcal{F} \mathcal{R} = \{\mathcal{R}_{\mathbf{x}}^T \partial^2 \mathcal{F}_{\mathbf{x}} \mathcal{R}_{\mathbf{x}}, \mathbf{x} \in \mathbb{R}^n\}$.

We prove the following lemma:

Lemma 2. $\partial^2\mathcal{F}'$ is a Gaussian random tensor field with autocovariance function $R_{\partial^2\mathcal{F}'}(\mathbf{s})$ such that

$$R_{\partial^2\mathcal{F}'}^0 = \sigma^2 \rho_0^{(4)} (\mathbf{I}_{m^2} + \mathbf{C}_m + \text{vec } \mathbf{I}_m \text{ vec}^\top \mathbf{I}_m). \quad (8)$$

Proof. Let $P_{\partial^2\mathcal{F}'|\mathcal{R}_x}$ be the conditional probability of the random tensor $\partial^2\mathcal{F}'_x$ given \mathcal{R}_x . Since $\partial^2\mathcal{F}_x$ is independent of \mathcal{R}_x , $\partial^2\mathcal{F}'_x$ given \mathcal{R}_x is a linear function of $\partial^2\mathcal{F}_x$, and therefore it is zero-mean Gaussian. Using the identity $\text{vec}(\mathbf{ABC}) = (\mathbf{C}^\top \otimes \mathbf{A}) \text{vec } \mathbf{B}$ and properties of commutator matrices [1] we can write $\text{vec } \partial^2\mathcal{F}'_x = (\mathcal{R}_x^\top \otimes \mathcal{R}_x^\top) \text{vec } \partial^2\mathcal{F}_x$, and therefore

$$R_{\partial^2\mathcal{F}'|\mathcal{R}_x}^0 = (\mathcal{R}_x^\top \otimes \mathcal{R}_x^\top) R_{\partial^2\mathcal{F}}^0 (\mathcal{R}_x \otimes \mathcal{R}_x) \quad (9)$$

$$= \sigma^2 \rho_0^{(4)} (\mathbf{I}_{m^2} + \mathbf{C}_m + \text{vec } \mathbf{I}_m \text{ vec}^\top \mathbf{I}_m), \quad (10)$$

using lemma 1. Since $\partial^2\mathcal{F}'_x$ given \mathcal{R}_x is zero-mean Gaussian and $R_{\partial^2\mathcal{F}'|\mathcal{R}_x}^0$ does not depend on \mathcal{R}_x for fixed m and n , we have $P_{\partial^2\mathcal{F}'|\mathcal{R}_x} = P_{\partial^2\mathcal{F}'}$. Therefore $P_{\partial^2\mathcal{F}'}$ is zero-mean Gaussian with autocovariance function satisfying (8). \square

Lemma 2 justifies the notation $\partial^2\mathcal{F}^m \triangleq \mathcal{R}^\top \partial^2\mathcal{F} \mathcal{R}$, for $\mathcal{R}(n, m)$. Let $\delta^2\mathcal{F}^n$ be the $(n(n+1)/2, 1)$ vector defined as $\delta^2\mathcal{F}^n \triangleq \text{uni } \partial^2\mathcal{F}^n = \mathbf{D}_n^+ \text{vec } \partial^2\mathcal{F}^n$. Its covariance matrix Σ_n is invertible and given by

$$\Sigma_n = \mathbf{D}_n^+ R_{\partial^2\mathcal{F}^n}^0 \mathbf{D}_n^{+, \top}. \quad (11)$$

Therefore the probability density $p_{\partial^2\mathcal{F}_x^n}$ of $\partial^2\mathcal{F}_x^n$ is standard:

$$p_{\delta^2\mathcal{F}_x^n}(\mathbf{h}) = \frac{1}{\sqrt{|2\pi\Sigma_n|}} \exp\left(-\frac{\mathbf{h}^\top \Sigma_n^{-1} \mathbf{h}}{2}\right). \quad (12)$$

We now define the random fields $(\mathcal{R}^n, \mathcal{L}^n) \triangleq \{(\mathcal{R}_x^n, \mathcal{L}_x^n) \in SO(n) \times \mathbb{R}^n, \mathbf{x} \in \mathbb{R}^n | \mathcal{R}_x^{n,\top} \mathcal{R}_x^n = \mathbf{I}_n \text{ and } \mathcal{R}_x^{n,\top} \text{diag}^{-1} \mathcal{L}_x^n \mathcal{R}_x^n = \partial^2\mathcal{F}_x^n\}$, where $SO(n)$ is the special orthogonal group of (n, n) matrices. Let eig_n^{-1} be the mapping given by

$$\begin{aligned} \text{eig}_n^{-1} : SO(n) \times \mathbb{R}^n &\rightarrow \mathbb{S}(n) \\ (\mathbf{R}, \boldsymbol{\lambda}) &\mapsto \mathbf{S} = \text{eig}_n^{-1}(\mathbf{R}, \boldsymbol{\lambda}), \end{aligned} \quad (13)$$

where $\mathbb{S}(n)$ is the set of (n, n) symmetric matrices. This mapping is differentiable and onto, and therefore the joint probability density $p_{\mathcal{R}_x, \mathcal{L}_x}$ of \mathcal{R}_x and \mathcal{L}_x is given by

$$p_{\mathcal{R}_x^n, \mathcal{L}_x^n}(\mathbf{R}, \boldsymbol{\lambda}) = p_{\partial^2\mathcal{F}_x^n}(\text{uni}(\mathbf{R}^\top \text{diag}^{-1} \boldsymbol{\lambda} \mathbf{R})) |J(\mathbf{R}, \boldsymbol{\lambda})|, \quad (14)$$

$|J(\mathbf{R}, \lambda)|$ is the absolute value of the Jacobian determinant of eig_n^{-1} .

Theorem 3.3.1 in [5] provides a “closed-form” expression of the probability density of the eigenvalues of random matrices in the *Gaussian orthogonal ensemble* (GOE_n). This is the ensemble of (n, n) real symmetric matrices \mathcal{M} with probability density invariant with respect to similarity transformations $\mathcal{M} \rightarrow \mathbf{R}^T \mathcal{M} \mathbf{R}$ for any given (n, n) orthonormal \mathbf{R} and such that the probability distribution of distinct entries are independent from each other. Even though each realization of $\partial^2 \mathcal{F}^n$ is real, symmetric, and, by applying lemma 2, invariant to the same similarity transformations, $\partial^2 \mathcal{F}^n$ is different from GOE_n , because the distinct entries of $\partial \mathcal{F}_x^n$ are not independent. However, the assumption of independency used in [5] is important only to derive an expression for the joint probability density of the entries of random matrices in GOE_n , and we already have that for the matrices in $\partial^2 \mathcal{F}^n$. Once such an expression is available the result in [5] derives from the observation that, in an expression analogous to (14), the term \mathbf{R} appeared only on $|J(\mathbf{R}, \lambda)|$, and therefore the probability density of the eigenvalues of matrices in GOE_n is obtained through the integration of $|J(\mathbf{R}, \lambda)|$ over $SO(n)$. The next lemma shows that this is also the case for the probability density $p_{\mathcal{L}_x^n}$ of the eigenvalues of $\partial^2 \mathcal{F}_x^n$:

Lemma 3. *Let $\lambda \in \mathbb{R}^n$, $\mathbf{R} \in SO(n)$, and Σ_n as in (11). Then*

$$p_{\partial^2 \mathcal{F}_x^n}(\text{uni}(\mathbf{R}^T \text{diag}^{-1} \lambda \mathbf{R})) = \frac{1}{\sqrt{|2\pi \Sigma_n|}} \exp\left(-\frac{\lambda^T \tilde{\Sigma}_n^{-1} \lambda}{2}\right), \quad (15)$$

where $\tilde{\Sigma}_n = \sigma^2 \rho_0^{(4)}(2\mathbf{I}_n + \mathbf{1}_{n,1} \mathbf{1}_{n,1}^T)$.

Proof. The following identity can be easily verified:

$$(R_{\partial^2 \mathcal{F}^n}^0)^+ = \frac{1}{4\sigma^2 \rho_0^{(4)}} \left(\mathbf{I}_{n^2} + \mathbf{C}_n - \frac{2}{2+n} \text{vec } \mathbf{I}_n \text{vec}^T \mathbf{I}_n \right). \quad (16)$$

Let $\Lambda = \text{diag}^{-1} \lambda$ and $\mathbf{u}_R \triangleq \text{uni}(\mathbf{R}^T \Lambda \mathbf{R})$. Therefore,

$$\begin{aligned} \mathbf{u}_R^T \Sigma_n^{-1} \mathbf{u}_R &= [\text{uni}(\mathbf{R}^T \Lambda \mathbf{R})]^T \Sigma_n^{-1} [\text{uni}(\mathbf{R}^T \Lambda \mathbf{R})] \\ &= [\mathbf{D}_n^+ \text{vec}(\mathbf{R}^T \Lambda \mathbf{R})]^T \Sigma_n^{-1} [\mathbf{D}_n^+ \text{vec}(\mathbf{R}^T \Lambda \mathbf{R})] \\ &= (\text{vec } \Lambda)^T (\mathbf{R} \otimes \mathbf{R}) \mathbf{D}_n^{+,T} \Sigma_n^{-1} \mathbf{D}_n^+ (\mathbf{R}^T \otimes \mathbf{R}^T) \text{vec } \Lambda, \end{aligned}$$

and, using $[(\mathbf{R} \otimes \mathbf{R}) \mathbf{D}_n]^+ = \mathbf{D}_n^+ (\mathbf{R}^T \otimes \mathbf{R}^T)$,

$$\mathbf{u}_R^T \Sigma_n^{-1} \mathbf{u}_R = (\text{vec } \Lambda)^T [(\mathbf{R} \otimes \mathbf{R}) \mathbf{D}_n \Sigma_n \mathbf{D}_n^T (\mathbf{R}^T \otimes \mathbf{R}^T)]^+ \text{vec } \Lambda,$$

which, using $\mathbf{D}_n \Sigma_n \mathbf{D}_n^T = R_{\partial^2 \mathcal{F}_n}^0$, yields

$$\begin{aligned}\mathbf{u}_R^T \Sigma_n^{-1} \mathbf{u}_R &= (\text{vec } \Lambda)^T [(\mathbf{R} \otimes \mathbf{R}) R_{\partial^2 \mathcal{F}_n}^0 (\mathbf{R}^T \otimes \mathbf{R}^T)]^+ \text{vec } \Lambda \\ &= (\text{vec } \Lambda)^T (R_{\partial^2 \mathcal{F}_n}^0)^+ \text{vec } \Lambda, \\ &= \lambda^T \tilde{\Sigma}_n^{-1} \lambda\end{aligned}\tag{17}$$

using lemma 1 and (16). \square

The integration of $|J(\mathbf{R}, \lambda)|$ over $SO(n)$, carried out in [5], gives

$$\int_{SO(n)} |J(\mathbf{R}, \lambda)| d\mathbf{R} = \left(\frac{\pi^{(n+1)/4}}{2}\right)^n \frac{\prod_{i=1}^{n-1} \prod_{j=i+1}^n |\lambda_j - \lambda_i|}{\prod_{i=1}^n \Gamma(1 + i/2)},\tag{18}$$

and it can be shown that the determinant of Σ_n is given by

$$|\Sigma_n| = 2^{n-1} (2+n) (\sigma^2 \rho_0^{(4)})^{n(n+1)/2}.\tag{19}$$

Together with lemma 3, these results demonstrate the following theorem:

Theorem 2. *The probability distribution $p_{\mathcal{L}_x^n}$ of the eigenvalues of $\partial^2 \mathcal{F}_x^n$ is*

$$\begin{aligned}p_{\mathcal{L}_x^n}(\lambda) &= \frac{2^{(2-7n-n^2)/4}}{\sqrt{2+n} (\sigma^2 \rho_0^{(4)})^{n(n+1)/4} \prod_{i=1}^n \Gamma(1+i/2)} \times \\ &\quad \prod_{i=1}^{n-1} \prod_{j=i+1}^n |\lambda_j - \lambda_i| \exp\left(-\frac{\lambda^T \tilde{\Sigma}_n^{-1} \lambda}{2}\right).\end{aligned}\tag{20}$$

Since \mathcal{N}_x in (6) is a function of $\partial \mathcal{F}_x^n$, theorem 1(1b) implies that $\partial^2 \mathcal{F}_x^n$ is independent of \mathcal{N}_x for all x . Therefore $\partial \mathcal{F}_x^{n-1} = \mathcal{N}_x \partial \mathcal{F}_x^n \mathcal{N}_x$ according to lemma 2. Theorem 2 can then be applied to obtain an expression for the probability density of the eigenvalues of the numerator of (6), $p_{\mathcal{L}_x^{n-1}}$. Using theorem 1(1a), we can show that the denominator of (6), $\|\partial \mathcal{F}_x^n\|$, is distributed according to $\sigma(-\rho_0^{(0)})^{1/2} \mathcal{X}(n)$, where $\mathcal{X}(n)$ follows a χ -distribution with n degrees of freedom, and therefore its probability density $p_{\|\partial \mathcal{F}_x^n\|}$ is given by

$$p_{\|\partial \mathcal{F}_x^n\|}(u) = \frac{2u^{n-1} \exp[u^2/(2\sigma^2 \rho_0^{(2)})]}{(-2\sigma^2 \rho_0^{(2)})^{n/2} \Gamma(n/2)},\tag{21}$$

We can now prove our main result:

Theorem 3. *Let \mathcal{K} be as in (6). Then*

$$\begin{aligned}p_{\mathcal{K}_x}(\kappa) &= \frac{2^{(n^2-7n+8)/4} \Gamma[n(n+1)/4]}{\sqrt{1+n} \Gamma(n/2) \prod_{i=1}^{n-1} \Gamma(1+i/2)} \times \\ &\quad \frac{\alpha^{n(n-1)/4} \prod_{i=1}^{n-2} \prod_{j=i+1}^{n-1} |\kappa_j - \kappa_i|}{\{\alpha [\sum_{i=1}^{n-1} \kappa_i^2 - \frac{1}{n+1} (\sum_{i=1}^{n-1} \kappa_i)^2] + 1\}^{\frac{n^2+n}{4}}},\end{aligned}\tag{22}$$

where $\alpha = -\rho_0^{(2)}/(2\rho_0^{(4)})$.

Proof. Since $\partial\mathcal{F}_x^n$ and $\partial^2\mathcal{F}_x^{n-1}$ are independent, so will be $\|\partial\mathcal{F}_x^n\|$ and \mathcal{L}_x^{n-1} . Using (6), we have.

$$p_{\mathcal{K}_x}(\kappa) = \int_0^\infty u^{n-1} p_{\mathcal{L}_x^{n-1}}(\kappa u) p_{\|\partial\mathcal{F}_x^n\|}(u) du. \quad (23)$$

Substituting (20) and (21) in (23), we obtain (22). \square

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APPENDIX: DIFFERENTIABILITY OF THE AUTOCORRELATION FUNCTION

Correlation functions are characterized by the Wiener-Khintchine theorem, a simplified version of which, shown below, is quoted verbatim from [2]:

Theorem 4. *A real function $\rho(\|\mathbf{s}\|)$ on \mathbb{R}^n is a correlation function if and only if it can be represented in the form*

$$\rho(\|\mathbf{s}\|) = 2^{(d-2)/2} \Gamma(d/2) \int_0^\infty \frac{J_{(d-2)/2}(ks)}{(ks)^{(d-2)/2}} d\Phi(k), \quad (\text{A.1})$$

where the function $\Phi(k)$ on \mathbb{R} has the properties of a distribution function and J_v is a Bessel function of the first kind and order v .

Lemma 4. *Let the i -th moment of the distribution $\Phi(k)$ in theorem 4 be defined. Then, the i -th derivative of $r(\|\mathbf{s}\|)$, $r^{(i)}(\|\mathbf{s}\|)$, exists and is given by*

$$r^{(i)}(\|\mathbf{s}\|) = 2^{(d-2)/2} \Gamma(d/2) \int_0^\infty k^i \frac{J_{(d-2)/2}(ks)}{(ks)^{(d-2)/2}} d\Phi(k). \quad (\text{A.2})$$

Proof. Define the operator D_i acting on a function $f(u)$ as

$$D_i[f(u)] = \left(\frac{1}{u} \frac{d}{du} \right)_i [f(u)] \quad (\text{A.3})$$

where the term in the right-hand side is recursively defined as

$$\left(\frac{1}{u} \frac{d}{du} \right)_1 [f(u)] = \frac{1}{u} \frac{df(u)}{du} \quad (\text{A.4})$$

$$\left(\frac{1}{u} \frac{d}{du} \right)_i [f(u)] = \left(\frac{1}{u} \frac{d}{du} \right) \left[\left(\frac{1}{u} \frac{d}{du} \right)_{i-1} [f(u)] \right]. \quad (\text{A.5})$$

It can be shown by induction that

$$r^{(i)}(\|\mathbf{s}\|) = \left(\frac{1}{u} \frac{d}{du} \right)_i [\rho(u)] \Bigg|_{u=\sqrt{\|\mathbf{s}\|}} . \quad (\text{A.6})$$

The operator D_i and the integral in theorem 4 can be interchanged, since the functions and the measure $d\Phi(k)$ involved satisfy the conditions of Lebesgue's dominated convergence theorem.

The identity

$$D_i \left[\frac{J_\nu(u)}{u^\nu} \right] = (-1)^i \frac{J_{\nu+i}(u)}{u^{\nu+i}}, \quad (\text{A.7})$$

found in [6], completes the proof. \square

- [1] J. R. Magnus and H. Neudecker, *Matrix Differential Calculus with Applications in Statistics and Econometrics*, Wiley Series in Probability and Mathematical Statistics (John Wiley & Sons, New York, 1995).
- [2] P. Abrahamsen, Tech. Rep. 917, Norwegian Computing Center, Oslo, Norway (1997), URL http://publications.nr.no/917_Rapport.pdf.
- [3] B. O'Neill, *Elementary Differential Geometry* (Academic Press, New York, 1966).
- [4] M. Spivak, *A Comprehensive Introduction to Differential Geometry*, vol. III (Publish or Perish, Houston, TX, USA, 1999), 3rd ed.
- [5] M. L. Mehta, *Random Matrices*, no. 142 in Pure and Applied Mathematics Series (Elsevier, San Diego, CA, 2004).
- [6] M. Abramowitz and I. A. Stegun, eds., *Handbook of Mathematical Functions with Formulas, Graphs, and Mathematical Tables*, vol. 55 of Applied Mathematics Series (National Bureau of Standards, Washington, D.C., USA, 1972), 10th ed.